

STAT 666 FORMULA SHEET

$$\bar{\mathbf{x}} = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \\ \vdots \\ \bar{x}_p \end{bmatrix} = \begin{bmatrix} \frac{1}{n} \sum_{i=1}^n x_{i1} \\ \frac{1}{n} \sum_{i=1}^n x_{i2} \\ \vdots \\ \frac{1}{n} \sum_{i=1}^n x_{ip} \end{bmatrix} = \frac{1}{n} \mathbf{X}' \mathbf{1}$$

$$\mathbf{S} = \begin{bmatrix} s_{11} & \cdots & s_{1p} \\ s_{21} & \cdots & s_{2p} \\ \vdots & & \vdots \\ s_{p1} & \cdots & s_{pp} \end{bmatrix} = \frac{1}{n-1} \mathbf{X}' (\mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}') \mathbf{X},$$

$$\mathbf{D} = \text{diag}(s_{11}, s_{22}, \dots, s_{pp})$$

$$\mathbf{D}^{-1/2} = \begin{bmatrix} \left(\frac{1}{s_{11}}\right)^{1/2} & & \mathbf{0} \\ & \ddots & \\ \mathbf{0} & & \left(\frac{1}{s_{pp}}\right)^{1/2} \end{bmatrix}.$$

$$\mathbf{R} = \mathbf{D}^{-1/2} \mathbf{S} \mathbf{D}^{-1/2}$$

$$\mathbf{A}_{(m+q) \times (n+r)} = \begin{bmatrix} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{A}_{22} \end{bmatrix}_{\substack{m \times n & m \times r \\ q \times n & q \times r}} \mathbf{B}_{(n+r) \times (s+t)} = \begin{bmatrix} \mathbf{B}_{11} & \mathbf{B}_{12} \\ \mathbf{B}_{21} & \mathbf{B}_{22} \end{bmatrix}_{\substack{n \times s & n \times t \\ r \times s & r \times t}}.$$

– Since all of the submatrices are conformable,

$$\mathbf{AB} = \begin{bmatrix} \mathbf{A}_{11}\mathbf{B}_{11} + \mathbf{A}_{12}\mathbf{B}_{21} & \mathbf{A}_{11}\mathbf{B}_{12} + \mathbf{A}_{12}\mathbf{B}_{22} \\ \mathbf{A}_{21}\mathbf{B}_{11} + \mathbf{A}_{22}\mathbf{B}_{21} & \mathbf{A}_{21}\mathbf{B}_{12} + \mathbf{A}_{22}\mathbf{B}_{22} \end{bmatrix}$$

– If $\mathbf{A}_{11}$ and $\mathbf{A}_{22}$ are square ($m = n$ and $q = r$) and $\mathbf{A}_{12} = \mathbf{A}_{21}' = \mathbf{0}$,

$$|\mathbf{A}| = |\mathbf{A}_{11}| |\mathbf{A}_{22}|$$

– If $\mathbf{A}_{11}$ and $\mathbf{A}_{22}$ are both square and nonsingular,

$$\begin{aligned} |\mathbf{A}| &= |\mathbf{A}_{11}| |\mathbf{A}_{22} - \mathbf{A}_{21} \mathbf{A}_{11}^{-1} \mathbf{A}_{12}| \\ &= |\mathbf{A}_{22}| |\mathbf{A}_{11} - \mathbf{A}_{12} \mathbf{A}_{22}^{-1} \mathbf{A}_{21}| \end{aligned}$$

$$\underbrace{|\mathbf{A} - \lambda \mathbf{I}_p|}_{\text{“characteristic equation”}} = 0$$

$$\mathbf{A} = \mathbf{CDC}'$$

$$\text{tr}(\mathbf{A}) = \sum_{i=1}^p \lambda_i$$

$$|\mathbf{A}| = \prod_{i=1}^p \lambda_i$$

$$\text{var}\{\bar{\mathbf{x}}\} = \frac{1}{n} \mathbf{\Sigma}$$

$$f_{\mathbf{x}}(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})' \mathbf{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})}$$

Constant probability density contour = $\{\text{all } \mathbf{x} \ni (\mathbf{x} - \boldsymbol{\mu})' \mathbf{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) = c^2\}$

If $\mathbf{x}_{p \times 1} \sim N_p(\boldsymbol{\mu}, \mathbf{\Sigma})$ and $\begin{pmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \end{pmatrix} \sim N_p\left(\begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix}, \begin{bmatrix} \mathbf{\Sigma}_{11} & \mathbf{\Sigma}_{12} \\ \mathbf{\Sigma}_{21} & \mathbf{\Sigma}_{22} \end{bmatrix}\right)$

$$\bullet \mathbf{Ax} + \mathbf{c} \sim N_q(\mathbf{A}\boldsymbol{\mu} + \mathbf{c}, \mathbf{A}\mathbf{\Sigma}\mathbf{A}')$$

$$* \mathbf{c}'\mathbf{x} \sim N_1(\mathbf{c}'\boldsymbol{\mu}, \mathbf{c}'\mathbf{\Sigma}\mathbf{c})$$

$$* E\{\mathbf{\Sigma}^{-\frac{1}{2}}(\mathbf{x} - \boldsymbol{\mu})\} = \mathbf{0}_p$$

$$\text{var}\{\mathbf{\Sigma}^{-\frac{1}{2}}(\mathbf{x} - \boldsymbol{\mu})\} = \mathbf{I}_p$$

$$\text{and } \mathbf{\Sigma}^{-\frac{1}{2}}(\mathbf{x} - \boldsymbol{\mu}) \sim N_p(\mathbf{0}, \mathbf{I})$$

$$\bullet \mathbf{x}_1 | \mathbf{x}_2 \sim N(\boldsymbol{\mu}_1 + \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22}^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2), \mathbf{\Sigma}_{11} - \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22}^{-1} \mathbf{\Sigma}_{21})$$

$$\bullet (\mathbf{x} - \boldsymbol{\mu})' \mathbf{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \sim \chi_p^2$$

Normal QQ Plot: Plot $\Phi^{-1}\left(\frac{i-\frac{1}{2}}{n}\right)$ vs. $x_{(1)}, \dots, x_{(n)}$

χ^2 QQ Plot: Plot $\left(\frac{i-\frac{1}{2}}{n}\right)^{th}$ quantile of χ_p^2 vs. $D_{(i)}^2 = i^{th}$ ordered value of $D_i^2 = (\mathbf{x}_i - \bar{\mathbf{x}})' \mathbf{S}^{-1}(\mathbf{x}_i - \bar{\mathbf{x}})$

Box and Cox (1964) recommend using

$$x^{(\lambda)} = \begin{cases} \frac{x^\lambda - 1}{\lambda} & \text{for } \lambda \neq 0 \\ \ln(x) & \text{for } \lambda = 0 \end{cases}$$

where λ is chosen by maximizing

$$\ell(\lambda) = -\frac{n}{2} \ln \left[\frac{1}{n} \sum_{i=1}^n (x_i^{(\lambda)} - \widetilde{x^{(\lambda)}})^2 \right] + (\lambda - 1) \sum_{i=1}^n \ln(x_i)$$

$$\text{and } \widetilde{x^{(\lambda)}} = \frac{1}{n} \sum_{i=1}^n x_i^{(\lambda)}$$

CLT: Let $\mathbf{x}_1, \dots, \mathbf{x}_n$ be independent obs. from a population with mean $\boldsymbol{\mu}$ and variance $\mathbf{\Sigma}$.

$$\bullet \sqrt{n}(\bar{\mathbf{x}} - \boldsymbol{\mu}) \text{ is approx. } N_p(\mathbf{0}, \mathbf{\Sigma})$$

$$\bullet n(\bar{\mathbf{x}} - \boldsymbol{\mu})' \mathbf{S}^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}) \text{ is approx } \chi_p^2 \text{ for } n - p \text{ large.}$$

Let $\mathbf{x}_i = \begin{bmatrix} \mathbf{x}_i^{(1)} \\ \mathbf{x}_i^{(2)} \end{bmatrix}$ $\leftarrow$ missing components $(q \times 1)$
 $\leftarrow$ observed components $((p-q) \times 1)$

$$\tilde{\boldsymbol{\mu}}_{p \times 1} = \begin{bmatrix} \tilde{\boldsymbol{\mu}}^{(1)} \\ \tilde{\boldsymbol{\mu}}^{(2)} \end{bmatrix} \text{ and } \tilde{\boldsymbol{\Sigma}} = \begin{bmatrix} \tilde{\boldsymbol{\Sigma}}_{11} & \tilde{\boldsymbol{\Sigma}}_{12} \\ \tilde{\boldsymbol{\Sigma}}_{21} & \tilde{\boldsymbol{\Sigma}}_{22} \end{bmatrix}$$

Each “E” step estimates $\mathbf{x}_i^{(1)}$ using regression:

$$\tilde{\mathbf{x}}_i^{(1)} = \tilde{\boldsymbol{\mu}}_i^{(1)} + \mathbf{B}_{q \times (p-q)} \underbrace{\left(\mathbf{x}_i^{(2)} - \tilde{\boldsymbol{\mu}}_i^{(2)} \right)}_{(p-q) \times 1}$$

where regression coefficients

$$\mathbf{B} = \begin{bmatrix} \tilde{\boldsymbol{\Sigma}}_{12} & \tilde{\boldsymbol{\Sigma}}_{22}^{-1} \\ q \times (p-q) & (p-q) \times (p-q) \end{bmatrix}$$

$$T^2 = n(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)' \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0) \sim T_{p,\nu}^2$$

$$\frac{\nu - p + 1}{\nu p} T_{p,\nu}^2 = F_{p,\nu-p+1}$$

100(1 - α)% C.R. for $\boldsymbol{\mu}$:

$$\{ \text{all } \boldsymbol{\mu} \ni \underbrace{n(\bar{\mathbf{x}} - \boldsymbol{\mu})' \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu})}_{\text{squared mult. distance from } \bar{\mathbf{x}}} \leq T_{\alpha,p,\nu}^2 \}$$

$$\mathbf{a}' \bar{\mathbf{x}} \pm t_{\frac{\alpha}{2}, n-1} \sqrt{\frac{\mathbf{a}' \mathbf{S} \mathbf{a}}{n}}$$

$$\mathbf{a}' \bar{\mathbf{x}} \pm \sqrt{T_{\alpha,p,\nu}^2} \sqrt{\frac{\mathbf{a}' \mathbf{S} \mathbf{a}}{n}}$$

$$\mathbf{a}' \bar{\mathbf{x}} \pm t_{\frac{\alpha}{2k}, n-1} \sqrt{\frac{\mathbf{a}' \mathbf{S} \mathbf{a}}{n}}$$

$$\mathbf{a} = \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu})$$

$$\mathbf{a}^* = \mathbf{D}^{\frac{1}{2}} \mathbf{a} \quad \text{where } \mathbf{D} = \begin{bmatrix} s_{11} & & \mathbf{0} \\ & \ddots & \\ \mathbf{0} & & s_{pp} \end{bmatrix}$$

$$T^2 = n(\mathbf{C}\bar{\mathbf{x}})'(\mathbf{CSC}')^{-1} \mathbf{C}\bar{\mathbf{x}} \sim T_{p-1, n-1}^2 \text{ (assuming } p-1 \text{ rows in } \mathbf{C})$$

$$T^2 = [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mathbf{S}_{p\ell} \right]^{-1} [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2] \sim T_{p, n_1 + n_2 - 2}^2$$

$$\mathbf{a}' \bar{\mathbf{x}}_1 - \mathbf{a}' \bar{\mathbf{x}}_2 \pm t_{\frac{\alpha}{2}, n_1 + n_2 - 2} \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mathbf{a}' \mathbf{S}_{p\ell} \mathbf{a}}$$

$$\mathbf{a}' \bar{\mathbf{x}}_1 - \mathbf{a}' \bar{\mathbf{x}}_2 \pm t_{\frac{\alpha}{2k}, n_1 + n_2 - 2} \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mathbf{a}' \mathbf{S}_{p\ell} \mathbf{a}}$$

$$\mathbf{a}' \bar{\mathbf{x}}_1 - \mathbf{a}' \bar{\mathbf{x}}_2 \pm \sqrt{T_{\alpha,p,\nu}^2} \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mathbf{a}' \mathbf{S}_{p\ell} \mathbf{a}}$$

$$\mathbf{a} = \mathbf{S}_{p\ell}^{-1} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)$$

$$\mathbf{a}^* = D_{p\ell}^{\frac{1}{2}} \mathbf{a}$$

$$\mathbf{D}_{p\ell} = \text{diag}\{\mathbf{S}_{p\ell}\} = \begin{bmatrix} s_{11,p\ell} & & \mathbf{0} \\ & \ddots & \\ \mathbf{0} & & s_{pp,p\ell} \end{bmatrix}$$

Nel and Van der Merwe (1986) test uses

$$T^{*2} = (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_e^{-1} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)$$

$$\mathbf{S}_e = \frac{\mathbf{S}_1}{n_1} + \frac{\mathbf{S}_2}{n_2}$$

$$T^{*2} \underset{\sim}{\text{approx}} T_{p,\nu^*}^2$$

$$\nu^* = \frac{\text{tr}\{\mathbf{S}_e^2\} + (\text{tr}\{\mathbf{S}_e\})^2}{\sum_{i=1}^2 \frac{1}{n_i - 1} \left[\text{tr}\left\{ \left(\frac{\mathbf{S}_i}{n_i} \right)^2 \right\} + \left(\text{tr}\left\{ \frac{\mathbf{S}_i}{n_i} \right\} \right)^2 \right]}$$

$$T_{\text{add}}^2 = (\nu - p) \frac{T_{p+q}^2 - T_p^2}{\nu + T_p^2} \sim T_{q,\nu-p}^2$$

$$\underbrace{\sum_{\ell=1}^g \sum_{j=1}^{n_\ell} (\mathbf{x}_{\ell_j} - \bar{\mathbf{x}})(\mathbf{\parallel})'}_{\text{total corrected sum of squares and cross products matrix}} = \underbrace{\sum_{\ell=1}^g n_\ell (\bar{\mathbf{x}}_\ell - \bar{\mathbf{x}})(\mathbf{\parallel})'}_{=\mathbf{H} \text{ “Between” groups matrix}} + \underbrace{\sum_{\ell=1}^g \sum_{j=1}^{n_\ell} (\mathbf{x}_{\ell_j} - \bar{\mathbf{x}}_\ell)(\mathbf{\parallel})'}_{=\mathbf{E} \text{ “Within” groups matrix}} = \sum_{\ell=1}^g (n_\ell - 1) \mathbf{S}_\ell$$

$$\Lambda = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}|} \sim \Lambda_{\alpha,p,\nu_H,\nu_E}$$

Let $\lambda_1, \dots, \lambda_s$ be the s non-zero eigenvalues of $\mathbf{E}^{-1} \mathbf{H}$, where $s = \min(p, \nu_H)$. Then $\Lambda = \prod_{i=1}^s \frac{1}{1 + \lambda_i}$

$$F = \frac{1 - \Lambda^{1/t}}{\Lambda^{1/t}} \frac{df_2}{df_1} \underset{\sim}{\text{approx}} F_{df_1, df_2}$$

$$df_1 = p\nu_H$$

$$df_2 = wt - \frac{1}{2}(p\nu_H - 2)$$

$$w = \nu_E + \nu_H - \frac{1}{2}(p + \nu_H + 1)$$

$$t = \begin{cases} \sqrt{\frac{p^2 \nu_H^2 - 4}{p^2 + \nu_H^2 - 5}} & \text{for } p^2 + \nu_H^2 - 5 > 0 \\ 1 & \text{for } p^2 + \nu_H^2 - 5 \leq 0 \end{cases}$$

$$\theta = \lambda_1$$

$$V = \sum_{i=1}^s \frac{\lambda_i}{1 + \lambda_i}$$

$$= \text{tr} \left\{ (\mathbf{E} + \mathbf{H})^{-1} \mathbf{H} \right\} = \sum_{i=1}^s \xi_i$$

where $\xi_1, \dots, \xi_s$ are the s ordered e'vals of $(\mathbf{E} + \mathbf{H})^{-1} \mathbf{H}$

$$\xi_i = \frac{\lambda_i}{1 + \lambda_i} \text{ and } \lambda_i = \frac{\xi_i}{1 - \xi_i}$$

$$U = \sum_{i=1}^s \lambda_i$$

$$= \text{tr} \{ \mathbf{E}^{-1} \mathbf{H} \}$$

$$\hat{\boldsymbol{\delta}} = c_1 \bar{\mathbf{x}}_1 + c_2 \bar{\mathbf{x}}_2 + \dots + c_g \bar{\mathbf{x}}_g$$

$$\text{var}\{\hat{\boldsymbol{\delta}}\} = \sum_{i=1}^g c_i^2 \frac{\boldsymbol{\Sigma}}{n_i} = \left(\sum_{i=1}^g \frac{c_i^2}{n_i} \right) \boldsymbol{\Sigma}$$

$$\widehat{\text{var}}\{\hat{\boldsymbol{\delta}}\} = \left(\sum_{i=1}^g \frac{c_i^2}{n_i} \right) \mathbf{S}_{p\ell}$$

where $\mathbf{S}_{p\ell} = \frac{1}{\nu_E} \mathbf{E}$ and $\nu_E = \sum_{i=1}^g (n_i - 1)$.

$$T^2 = \hat{\boldsymbol{\delta}}' \left[\widehat{\text{var}} \left\{ \hat{\boldsymbol{\delta}} \right\} \right]^{-1} \hat{\boldsymbol{\delta}} \sim T_{p, \nu_E}^2$$

$$\mathbf{H}_1 = \frac{1}{\sum_{i=1}^g \frac{c_i^2}{n_i}} \hat{\boldsymbol{\delta}} \hat{\boldsymbol{\delta}}' \text{ and } \Lambda = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}_1|} \sim \Lambda_{p, 1, \nu_E}$$

100(1 - α)% (Simultaneous) Confidence Interval for $\mu_{ij} - \mu_{kj}$ is:

$$(\bar{x}_{ij} - \bar{x}_{kj}) \pm t \underbrace{\left[\frac{\alpha/2}{\text{# of comparisons}} \right]}_{\substack{\frac{\alpha}{2} \text{ divided by} \\ \text{# of comparisons} \\ = pg(g-1)/2}}, \left[\sum_{i=1}^g (n_i - 1) \right] \sqrt{s_{p\ell, jj} \left(\frac{1}{n_i} + \frac{1}{n_k} \right)}$$

where $s_{p\ell, jj}$ is the j^{th} diagonal element of $\mathbf{S}_{p\ell} = \mathbf{E} / (\sum_{i=1}^g (n_i - 1))$

λ_1 = largest e'value of $\mathbf{E}^{-1} \mathbf{H}$ and $\mathbf{a}_1$ is corresp. e'vec

$$a_i^* = a_i \sqrt{s_{p\ell, ii}} = a_i \times \sqrt{\frac{1}{\nu_E} e_{ii}}$$

$$\Lambda_{z|y} = \frac{\Lambda_{yz}}{\Lambda_y} \sim \Lambda_{q, \nu_H, \nu_E - p}$$

$\uparrow$ "partial Λ statistic" $\uparrow$ # of vars in $\mathbf{z}$ $\uparrow$ # of vars in $\mathbf{y}$

$$\Lambda_A = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}_A|} \sim \Lambda_{p, a-1, ab(n-1)}$$

$$\Lambda_B = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}_B|} \sim \Lambda_{p, b-1, ab(n-1)}$$

$$\Lambda_{AB} = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}_{AB}|} \sim \Lambda_{p, (a-1)(b-1), ab(n-1)}$$

Test for H_{01} :

$$\Lambda = \frac{|\mathbf{C} \mathbf{E} \mathbf{C}'|}{|\mathbf{C}(\mathbf{E} + \mathbf{H}) \mathbf{C}'|} \sim \Lambda_{p-1, \nu_H, \nu_E}$$

Test for H_{02} :

$$\Lambda = \frac{\mathbf{1}' \mathbf{E} \mathbf{1}}{\mathbf{1}' \mathbf{E} \mathbf{1} + \mathbf{1}' \mathbf{H} \mathbf{1}} \sim \Lambda_{1, \nu_H, \nu_E}$$

$$\Rightarrow \frac{1 - \Lambda}{\Lambda} \frac{\nu_E}{\nu_H} \sim F_{\nu_H, \nu_E}$$

Test for H_{03} :

$$T^2 = \left(\sum_{\ell=1}^g n_{\ell} \right) (\mathbf{C} \bar{\mathbf{x}})' \left(\frac{1}{\nu_e} \mathbf{C} \mathbf{E} \mathbf{C}' \right)^{-1} \mathbf{C} \bar{\mathbf{x}} \sim T_{p-1, \nu_E}^2$$

$$\Rightarrow \frac{\nu_E - (p-1) + 1}{\nu_E (p-1)} T^2 \sim F_{p-1, \nu_E - (p-1) + 1}$$

$$\varepsilon = \frac{\left[\text{tr} \left(\boldsymbol{\Sigma} - \frac{1}{p} \mathbf{1} \mathbf{1}' \boldsymbol{\Sigma} \right) \right]^2}{(p-1) \text{tr} \left[\left(\boldsymbol{\Sigma} - \frac{1}{p} \mathbf{1} \mathbf{1}' \boldsymbol{\Sigma} \right)^2 \right]}$$

SAS: "G - G ε "

Test for A : $F = \frac{MSA}{MSE} \sim F_{\hat{\varepsilon}(p-1), \hat{\varepsilon}g(n-1)(p-1)}$

Test for AB : $F = \frac{MSAB}{MSE} \sim F_{\hat{\varepsilon}(g-1)(p-1), \hat{\varepsilon}g(n-1)(p-1)}$

$$T^2 = \underbrace{N}_{\sum_{i=1}^g n_i} \underbrace{(\mathbf{C} \bar{\mathbf{x}})'}_{\substack{\uparrow \\ \text{grand mean}}} (\underbrace{\mathbf{C} \mathbf{S}_{p\ell} \mathbf{C}'}_{\substack{\uparrow \\ \mathbf{E} \\ \nu_E}})^{-1} (\underbrace{\mathbf{C} \bar{\mathbf{x}}}_{(p-1) \times p}) \sim T_{p-1, \nu_E}^2$$

$$\Lambda = \frac{\mathbf{1}' \mathbf{E} \mathbf{1}}{\mathbf{1}' \mathbf{E} \mathbf{1} + \mathbf{1}' \mathbf{H} \mathbf{1}} \sim \Lambda_{1, \nu_H, \nu_E}$$

$$\Lambda = \frac{|\mathbf{C} \mathbf{E} \mathbf{C}'|}{|\mathbf{C}(\mathbf{E} + \mathbf{H}) \mathbf{C}'|} \sim \Lambda_{p-1, \nu_H, \nu_E}$$

A and B are within subjects factors and C is a between subjects factor

Test for A :

$$T^2 = N(\mathbf{A}\bar{\mathbf{x}})'(\mathbf{A}\mathbf{S}_{p\ell}\mathbf{A}')^{-1}\mathbf{A}\bar{\mathbf{x}} \sim T_{a-1, \nu_E}^2$$

$\uparrow$ grand mean $\uparrow$ $\frac{1}{\nu_E}\mathbf{E}$ $\uparrow$ $\sum_{i=1}^g(n_i - 1)$
 when only one between subjects factor is used

OR

$$\Lambda = \frac{|\mathbf{A}\mathbf{E}\mathbf{A}'|}{|\mathbf{A}(\mathbf{E} + \mathbf{H}^*)\mathbf{A}'|} \sim \Lambda_{a-1, 1, \nu_E}$$

where $\mathbf{H}^* = N\bar{\mathbf{x}}\bar{\mathbf{x}}'$ and $\mathbf{A} = \mathbf{A}^* \otimes \mathbf{1}'_b$

Test for B :

$$T^2 = N(\mathbf{B}\bar{\mathbf{x}})'(\mathbf{B}\mathbf{S}_{p\ell}\mathbf{B}')^{-1}\mathbf{B}\bar{\mathbf{x}} \sim T_{b-1, \nu_E}^2$$

OR

$$\Lambda = \frac{|\mathbf{B}\mathbf{E}\mathbf{B}'|}{|\mathbf{B}(\mathbf{E} + \mathbf{H}^*)\mathbf{B}'|} \sim \Lambda_{b-1, 1, \nu_E}$$

where $\mathbf{B} = \mathbf{1}'_a \otimes \mathbf{B}^*$

Test of AB interaction:

$$T^2 = N(\mathbf{G}\bar{\mathbf{x}})'(\mathbf{G}\mathbf{S}_{p\ell}\mathbf{G}')^{-1}\mathbf{G}\bar{\mathbf{x}} \sim T_{(a-1)(b-1), \nu_E}^2$$

OR

$$\Lambda = \frac{|\mathbf{G}\mathbf{E}\mathbf{G}'|}{|\mathbf{G}(\mathbf{E} + \mathbf{H}^*)\mathbf{G}'|} \sim \Lambda_{(a-1)(b-1), 1, \nu_E}$$

where $\mathbf{G} = \mathbf{A}^* \otimes \mathbf{B}^*$

Test for C :

$$\Lambda = \frac{\mathbf{1}'\mathbf{E}\mathbf{1}}{\mathbf{1}'\mathbf{E}\mathbf{1} + \mathbf{1}'\mathbf{H}\mathbf{1}} \sim \Lambda_{1, \nu_H, \nu_E}$$

Tests for AC, BC, ABC interactions:

$$\Lambda = \frac{|\mathbf{A}\mathbf{E}\mathbf{A}'|}{|\mathbf{A}(\mathbf{E} + \mathbf{H})\mathbf{A}'|} \sim \Lambda_{a-1, \nu_H, \nu_E}$$

$$\Lambda = \frac{|\mathbf{B}\mathbf{E}\mathbf{B}'|}{|\mathbf{B}(\mathbf{E} + \mathbf{H})\mathbf{B}'|} \sim \Lambda_{b-1, \nu_H, \nu_E}$$

$$\Lambda = \frac{|\mathbf{G}\mathbf{E}\mathbf{G}'|}{|\mathbf{G}(\mathbf{E} + \mathbf{H})\mathbf{G}'|} \sim \Lambda_{(a-1)(b-1), \nu_H, \nu_E}$$

$$\begin{aligned} \text{vec } \mathbf{Y} &= \text{vec } (\mathbf{X}\mathbf{B}) + \text{vec } (\mathbf{\Xi}) \\ &= \underbrace{(\mathbf{I}_p \otimes \mathbf{X})}_{\substack{\uparrow \\ \text{rank} = pr \\ \text{when rank}(\mathbf{X})=r}} \underbrace{\text{vec } \mathbf{B}}_{\substack{\equiv \boldsymbol{\beta} \\ pr \times 1}} + \underbrace{\text{vec } (\mathbf{\Xi})}_{\substack{\equiv \mathbf{e} \\ np \times 1}} \end{aligned}$$

$$\underbrace{\mathbf{Y}'\mathbf{Y} - n\bar{\mathbf{y}}\bar{\mathbf{y}}'}_{\text{corrected total SS \& CP}} = \underbrace{\mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y}}_{=\mathbf{E}_{p \times p}} + \underbrace{\hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y} - n\bar{\mathbf{y}}\bar{\mathbf{y}}'}_{=\mathbf{H}_{p \times p}}$$

$$\Lambda = \frac{|\mathbf{E}|}{|\mathbf{E} + \mathbf{H}|} = \frac{|\mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y}|}{|\mathbf{Y}'\mathbf{Y} - n\bar{\mathbf{y}}\bar{\mathbf{y}}'|} \sim \Lambda_{p, \underbrace{q}_{r-1}, \underbrace{n-q-1}_{n-r}}$$

$$s = \min(p, q)$$

$$m = \frac{1}{2}(|q - p| - 1)$$

$$N = \frac{1}{2}(n - q - p - 2)$$

$$\frac{|\mathbf{S}|}{|\mathbf{S}_{xx}||\mathbf{S}_{yy}|}$$

$$\mathbf{H}_{\text{diff}} = \hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y} - \hat{\mathbf{B}}'_r\mathbf{X}'_r\mathbf{Y}$$

$$\mathbf{E}_{\text{full}} = \mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y}$$

$$\begin{aligned} \Lambda_{x_{q-h+1}, \dots, x_q | x_1, \dots, x_{q-h}} &= \frac{|\mathbf{E}_{\text{full}}|}{|\mathbf{E}_{\text{full}} + \mathbf{H}_{\text{diff}}|} = \frac{|\mathbf{E}_{\text{full}}|}{|\mathbf{E}_{\text{red}}|} \\ &= \frac{|\mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'\mathbf{X}'\mathbf{Y}|}{|\mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'_{\text{red}}\mathbf{X}'_{\text{red}}\mathbf{Y}|} \\ &= \frac{\Lambda_{\text{full}}}{\Lambda_{\text{red}}} \\ &\sim \Lambda_{p, h, n-q-1} \\ &\quad \uparrow \\ &\quad \# \text{ of } xs \end{aligned}$$

$$s = \min(p, h)$$

$$m = \frac{1}{2}(|h - p| - 1)$$

$$N = \frac{1}{2}(n - h - p - 2)$$

$$\hat{\mathbf{B}}_i = \begin{bmatrix} \hat{\beta}_{01} & \hat{\beta}_{02} & \cdots & \hat{\beta}_{0p} \\ \hat{\beta}_{i1} & \hat{\beta}_{i2} & \cdots & \hat{\beta}_{ip} \end{bmatrix}$$

$$\Lambda_{x_i} = \frac{|\mathbf{Y}'\mathbf{Y} - \hat{\mathbf{B}}'_i\mathbf{X}'\mathbf{Y}|}{|\mathbf{Y}'\mathbf{Y} - n\bar{\mathbf{y}}\bar{\mathbf{y}}'|} \sim \Lambda_{p, 1, n-2}$$

$$\Lambda_{x_i | x_1, \dots, x_j} \sim \Lambda_{p, 1, n-j-1}$$

$$\Lambda_{x_i | x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_m} \sim \Lambda_{p, 1, n-m-1}$$

$u_1 = \mathbf{a}'_1 \mathbf{y}$ and $v_1 = \mathbf{b}'_1 \mathbf{x}$ are the “first canonical variates”

$$(\mathbf{S}_{yy}^{-1} \mathbf{S}_{yx} \mathbf{S}_{xx}^{-1} \mathbf{S}_{xy} - r^2 \mathbf{I}_p) \mathbf{a} = \mathbf{0}$$

$$(\mathbf{S}_{xx}^{-1} \mathbf{S}_{xy} \mathbf{S}_{yy}^{-1} \mathbf{S}_{yx} - r^2 \mathbf{I}_q) \mathbf{b} = \mathbf{0}$$

$$\mathbf{c}_i = \begin{bmatrix} s_{y1} & 0 & \cdots & 0 \\ 0 & s_{y2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & s_{yp} \end{bmatrix} \mathbf{a}_i,$$

$$\mathbf{d}_i = \begin{bmatrix} s_{x1} & 0 & \cdots & 0 \\ 0 & s_{x2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & s_{xq} \end{bmatrix} \mathbf{b}_i$$

More simply, conduct analysis using $\mathbf{R}_{yy}^{-1} \mathbf{R}_{yx} \mathbf{R}_{xx}^{-1} \mathbf{R}_{xy}$ and $\mathbf{R}_{xx}^{-1} \mathbf{R}_{xy} \mathbf{R}_{yy}^{-1} \mathbf{R}_{yx}$ which have eigenvectors $\mathbf{c}_i$ and $\mathbf{d}_i$, respectively, and have eigenvalues $r_1^2, \dots, r_s^2$.

$$\Lambda = \prod_{i=1}^s (1 - r_i^2) \sim \Lambda_{p,q,n-1-q}$$

$$\Lambda_m = \prod_{i=m}^s (1 - r_i^2) \sim \Lambda_{p-m+1,q-m+1,n-m-q}$$

$$\text{or } \Lambda_{q-m+1,p-m+1,n-m-p}$$

$$[\text{since } \Lambda_{p,\nu_H,\nu_E} \stackrel{q}{=} \Lambda_{\nu_H,p,\nu_H+\nu_E-p}]$$

Test that the last k population eigenvalues (denoted $\lambda_{p-k+1}^{\text{pop}}, \dots, \lambda_p^{\text{pop}}$) are equal

$$H_0 : \lambda_{p-k+1}^{\text{pop}} = \dots = \lambda_p^{\text{pop}}$$

$$\bar{\lambda} = \frac{1}{k} \sum_{i=p-k+1}^p \lambda_i$$

$$u = \left(n - \frac{2p+11}{6} \right) \left(k \ln \bar{\lambda} - \sum_{i=p-k+1}^p \ln \lambda_i \right) \\ \approx \chi_{\frac{1}{2}(k-1)(k+2)}^2$$

$$\mathbf{x}_{p \times 1} = \underbrace{\boldsymbol{\mu}}_{p \times 1} + \underbrace{\mathbf{\Lambda}}_{p \times k} \underbrace{\mathbf{f}}_{k \times 1} + \underbrace{\mathbf{e}}_{p \times 1}$$

$$\begin{aligned} \text{var}\{\mathbf{x}\} &= \mathbf{\Lambda} \text{var}\{\mathbf{f}\} \mathbf{\Lambda}' + \text{var}\{\mathbf{e}\} \\ &= \mathbf{\Lambda} \mathbf{\Lambda}' + \mathbf{\Psi} \\ &= \mathbf{\Lambda} \mathbf{\Lambda}' + \mathbf{\Psi} \end{aligned}$$

$$\begin{aligned} \text{var}\{x_i\} &= [\lambda_{i1}, \lambda_{i2}, \dots, \lambda_{ik}] [\lambda_{i1}, \lambda_{i2}, \dots, \lambda_{ik}]' + \psi_i \\ \underbrace{\sigma_{ii}}_{\substack{\uparrow \\ \text{variance for} \\ \text{variable } i}} &= \underbrace{\lambda_{i1}^2 + \lambda_{i2}^2 + \dots + \lambda_{ik}^2}_{\substack{\uparrow \\ h_i^2 = \text{communality} \\ \text{for variable } i}} + \underbrace{\psi_i}_{\substack{\uparrow \\ \text{specific variance} \\ \text{for variable } i}} \end{aligned}$$

$$\mathbf{S} = \mathbf{A} \mathbf{D} \mathbf{A}'$$

$$\cong \mathbf{A}_1 \mathbf{D}_1 \mathbf{A}_1' \quad \left(\begin{array}{l} \text{retaining only } k < p \text{ factors,} \\ \text{where } \mathbf{A}_1 = [\mathbf{a}_1 \dots \mathbf{a}_k] \\ \text{and } \mathbf{D}_1 = \text{diag}(d_1 \dots d_k) \end{array} \right)$$

$$= (\mathbf{A}_1 \mathbf{D}_1^{1/2}) (\mathbf{A}_1 \mathbf{D}_1^{1/2})'$$

$$= \hat{\mathbf{\Lambda}} \hat{\mathbf{\Lambda}}'$$

$$\Rightarrow \hat{\mathbf{\Lambda}} = \mathbf{A}_1 \mathbf{D}_1^{1/2}$$

and

$$\hat{\Psi} = \mathbf{S} - \hat{\mathbf{\Lambda}} \hat{\mathbf{\Lambda}}' \quad \left(\hat{\psi}_i = s_{ii} - \underbrace{\sum_{j=1}^k \hat{\lambda}_{ij}^2}_{h_i^2 \text{ or "communality"}} \right)$$

$$\hat{\mathbf{\Lambda}} = \mathbf{A}_1^* \mathbf{D}_1^{*\frac{1}{2}}$$

where $\mathbf{A}^* \mathbf{D}^* \mathbf{A}^{*'} is the spectral decomposition of $\mathbf{S} - \mathbf{\Psi}^\circ$ (or $\mathbf{R} - \mathbf{\Psi}^\circ$)$

$$\left(n - \frac{2p+4k+11}{6} \right) \ln \left(\frac{|\hat{\mathbf{\Lambda}} \hat{\mathbf{\Lambda}} + \hat{\Psi}|}{|\mathbf{S}|} \right) \sim \chi_{\frac{1}{2}[(p-k)^2 - (p+k)]}^2$$

$$(p-k)^2 \text{ must be } \geq (p+k)$$

$$k \leq \frac{1}{2} (2p+1 - \sqrt{8p+1})$$

$$\hat{\mathbf{f}}_t = (\hat{\mathbf{\Lambda}}' \hat{\Psi}^{-1} \hat{\mathbf{\Lambda}}')^{-1} \hat{\mathbf{\Lambda}}' \hat{\Psi}^{-1} (\mathbf{x}_t - \bar{\mathbf{x}})$$

$$\hat{\mathbf{f}}_t = \hat{\mathbf{\Lambda}}' (\hat{\mathbf{\Lambda}} \hat{\mathbf{\Lambda}}' + \hat{\Psi})^{-1} (\mathbf{x}_t - \bar{\mathbf{x}})$$

$$\text{or } \hat{\mathbf{f}}_t = \hat{\mathbf{\Lambda}}' \mathbf{S}^{-1} (\mathbf{x}_t - \bar{\mathbf{x}})$$

$$\underbrace{\begin{bmatrix} \mathbf{x}_1 \\ \vdots \\ \mathbf{x}_{p-k} \end{bmatrix}}_{\mathbf{x}} = \underbrace{\begin{bmatrix} \boldsymbol{\beta}_0 \\ \mathbf{0} \end{bmatrix}}_{\boldsymbol{\mu}} + \underbrace{\begin{bmatrix} \mathbf{B} \\ \mathbf{I}_k \end{bmatrix}}_{\mathbf{\Lambda}} \mathbf{f} + \mathbf{e}$$

$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_{11} & \cdots & \lambda_{1k} \\ \vdots & & \\ \lambda_{(p-k),1} & \cdots & \lambda_{(p-k),k} \\ 1 & & 0 \\ & \ddots & \\ 0 & & 1 \end{bmatrix}$$

$$t = \frac{\hat{\lambda}_{21}}{\text{s.e.}(\hat{\lambda}_{21})}$$

Saturated model:

$$df = \frac{1}{2}[(p-k)^2 - (p+k)] \quad \leftarrow \boxed{df \text{ for a } \chi^2 \text{ GOF test}}$$

General model with maximum likelihood estimation:

$$\ell(\boldsymbol{\theta}, \mathbf{S}) = n \left(\log |\boldsymbol{\Sigma}[\boldsymbol{\theta}]| + \text{tr} \left\{ \mathbf{S} (\boldsymbol{\Sigma}[\boldsymbol{\theta}])^{-1} \right\} - \log |\mathbf{S}| - p \right)$$

$$\ell(\hat{\boldsymbol{\theta}}_{ML}; \mathbf{S}) \rightarrow \chi^2_{\frac{1}{2}p(p+1)-q}$$

is a goodness-of-fit statistic where q is # of params in model and $\frac{1}{2}p(p+1)$ is # of statistics in $\mathbf{S}$.

$$CFI = 1 - \frac{\max(\chi_M^2 - df_M, 0)}{\max(\chi^2 - df_B, \chi_M^2 - df_M, 0)}$$

$$RMSEA = \sqrt{\frac{1}{n-1} \frac{\max(\chi_M^2 - df_M, 0)}{df_M}}$$

$$SRMR = \sqrt{\frac{2}{p(p+1)} \sum_{i=1}^p \sum_{j=1}^i \frac{(s_{ij} - \sigma_{ij})^2}{s_{ii}s_{jj}}}$$

$$\mathbf{a} = \underset{\mathbf{b} \neq \mathbf{0}}{\operatorname{argmax}} \frac{[\mathbf{b}'(\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)]^2}{\mathbf{b}'\mathbf{S}_{p\ell}\mathbf{b}}$$

$$\mathbf{a}_i^* = \mathbf{D}^{1/2}\mathbf{a}_i, \quad i = 1, \dots, s$$

$$\text{where } \mathbf{D} = \frac{1}{\nu_E} \begin{bmatrix} e_{11} & & \mathbf{0} \\ & \ddots & \\ \mathbf{0} & & e_{pp} \end{bmatrix} \\ = \frac{1}{\nu_E} \text{diag}(\mathbf{E})$$

$$\text{ECM} = c\{2|1\} \cdot \Pr\{2|1\} \cdot p_1 + c\{1|2\} \cdot \Pr\{1|2\} \cdot p_2$$

ECM is minimized with the rule:

$$\hat{G}_1 = \{\mathbf{x} : p_1 f_1(\mathbf{x}) c\{2|1\} > p_2 f_2(\mathbf{x}) c\{1|2\}\}$$

and

$$\hat{G}_2 = \{\mathbf{x} : p_1 f_1(\mathbf{x}) c\{2|1\} < p_2 f_2(\mathbf{x}) c\{1|2\}\}$$

Assign $\mathbf{x}_0$ to $\hat{G}_1$ if:

$z_0 = \mathbf{a}'\mathbf{x}_0$ is closer to $\bar{z}_1 = \mathbf{a}'\bar{\mathbf{x}}_1$ than $\bar{z}_2 = \mathbf{a}'\bar{\mathbf{x}}_2$

or

$$z_0 > \frac{1}{2}(\bar{z}_1 + \bar{z}_2)$$

So that the rule is:

$$\hat{G}_1 = \left\{ \mathbf{x} : (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} \mathbf{x} > \frac{1}{2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} (\bar{\mathbf{x}}_1 + \bar{\mathbf{x}}_2) \right\}$$

and

$$\hat{G}_2 = \left\{ \mathbf{x} : (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} \mathbf{x} < \frac{1}{2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} (\bar{\mathbf{x}}_1 + \bar{\mathbf{x}}_2) \right\}$$

Incorporating costs and priors, rule is:

$$\hat{G}_1 = \left\{ \mathbf{x} : (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} \mathbf{x} > \frac{1}{2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} (\bar{\mathbf{x}}_1 + \bar{\mathbf{x}}_2) + \ln \frac{p_2 c\{1|2\}}{p_1 c\{2|1\}} \right\}$$

and

$$\hat{G}_2 = \left\{ \mathbf{x} : (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} \mathbf{x} < \frac{1}{2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_{p\ell}^{-1} (\bar{\mathbf{x}}_1 + \bar{\mathbf{x}}_2) + \ln \frac{p_2 c\{1|2\}}{p_1 c\{2|1\}} \right\}$$

Assign $\mathbf{x}$ to $\hat{G}_{i'}$ if:

$$L_i(\mathbf{x}) = \bar{\mathbf{x}}_i' \mathbf{S}_{p\ell}^{-1} \mathbf{x} - \frac{1}{2} \bar{\mathbf{x}}_i' \mathbf{S}_{p\ell}^{-1} \bar{\mathbf{x}}_i, \quad i = 1, \dots, g$$

is maximized when $i = i'$

For known priors, use:

$$L_i^*(\mathbf{x}) = \bar{\mathbf{x}}_i' \mathbf{S}_{p\ell}^{-1} \mathbf{x} - \frac{1}{2} \bar{\mathbf{x}}_i' \mathbf{S}_{p\ell}^{-1} \bar{\mathbf{x}}_i + \ln p_i, \quad i = 1, \dots, g$$

Quadratic functions:

$$D_i^2(\mathbf{x}) = (\mathbf{x} - \bar{\mathbf{x}}_i)' \mathbf{S}_i^{-1} (\mathbf{x} - \bar{\mathbf{x}}_i)$$

or

$$Q_i(\mathbf{x}) = \ln p_i - \frac{1}{2} \ln |\mathbf{S}_i| - \frac{1}{2} (\mathbf{x} - \bar{\mathbf{x}}_i)' \mathbf{S}_i^{-1} (\mathbf{x} - \bar{\mathbf{x}}_i)$$

Assign $\mathbf{x}$ to $\hat{G}_i$ if $k_i = \max_j k_j$

Rule assumes $p_i = \frac{n_i}{n}$, $i = 1, \dots, g$. If $p_i \neq \frac{n_i}{n}$, consider the revised rule:

$$\text{Assign } \mathbf{x} \text{ to } \hat{G}_i \text{ if } \frac{k_i p_i}{n_i} = \max_j \frac{k_j p_j}{n_j}$$

$$I_A = \sum_{i=1}^k p_{i|A} (1 - p_{i|A})$$

$$p_{i|A} = \frac{p_i(n_{iA}/n_i)}{\sum_{i=1}^k p_i(n_{iA}/n_i)}$$

$$p_A = \sum_{i=1}^k p_i(n_{iA}/n_i)$$

$$\Delta I = p_A I_A - (p_{A_L} I_{A_L} + p_{A_R} I_{A_R})$$

$$d(\mathbf{x}, \mathbf{y}) = \sqrt{(\mathbf{x} - \mathbf{y})'(\mathbf{x} - \mathbf{y})}$$

$$\checkmark \text{(a)} \quad (\mathbf{A} \otimes \mathbf{B})(\mathbf{C} \otimes \mathbf{D}) = (\mathbf{AC}) \otimes (\mathbf{BD})$$

$$\checkmark \text{(b)} \quad \text{vec}(\mathbf{ABC}) = (\mathbf{C}' \otimes \mathbf{A}) \text{vec} \mathbf{B}$$

$$\text{(c)} \quad \text{tr}\{\mathbf{AB}\} = (\text{vec} \mathbf{A}')' \text{vec} \mathbf{B} = (\text{vec} \mathbf{A})' \text{vec} \mathbf{B}'$$

$$\text{(d)} \quad \text{tr}\{\mathbf{ABCD}\} = (\text{vec} \mathbf{A}')' (\mathbf{D}' \otimes \mathbf{B}) \text{vec} \mathbf{C} = (\text{vec} \mathbf{A})' (\mathbf{B} \otimes \mathbf{D}') \text{vec} \mathbf{C}'$$

$$\checkmark \text{(e)} \quad (\mathbf{A} \otimes \mathbf{B})' = \mathbf{A}' \otimes \mathbf{B}'$$

$$\checkmark \text{(f)} \quad (\mathbf{A} \otimes \mathbf{B})^{-1} = \mathbf{A}^{-1} \otimes \mathbf{B}^{-1}$$